

Stat 511 Exam1

**February 24, 2009
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I have neither given nor received unauthorized assistance on this exam.

Name

Name Printed

1. Consider a segmented simple linear regression problem in one variable, x . In particular, suppose that $n = 6$ values of a response y are related to values $x = 0, 1, 2, 3, 4, 5$ by a Gauss-Markov normal linear model $\mathbf{Y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon}$ for

$$\mathbf{Y} = \begin{pmatrix} y_1 \\ y_2 \\ y_3 \\ y_4 \\ y_5 \\ y_6 \end{pmatrix}, \mathbf{X} = \begin{pmatrix} 1 & 0 & 0 \\ 1 & 1 & 0 \\ 1 & 2 & 0 \\ 1 & 3 & 1 \\ 1 & 4 & 2 \\ 1 & 5 & 3 \end{pmatrix}, \boldsymbol{\beta} = \begin{pmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \end{pmatrix}, \text{ and } \boldsymbol{\varepsilon} = \begin{pmatrix} \varepsilon_1 \\ \varepsilon_2 \\ \varepsilon_3 \\ \varepsilon_4 \\ \varepsilon_5 \\ \varepsilon_6 \end{pmatrix}$$

Values of x are in the second column of the model matrix. This model allows the linear form $y \approx \beta_0 + \beta_1 x$ for $x \leq 2$ and the linear form $y \approx \beta_0 + 2\beta_1 + (\beta_1 + \beta_2)(x - 2)$ for $x \geq 2$. Notice that there is continuity of these forms at $x = 2$.

10 pts

a) This a full rank model. **Argue** carefully that this is the case.

$$\text{Here } (\mathbf{X}'\mathbf{X})^{-1} = \begin{pmatrix} .825 & -.474 & .526 \\ -.474 & .421 & -.579 \\ .526 & -.579 & .921 \end{pmatrix} \text{ and for } \mathbf{Y}' = (0, 2, 4, 3, 1, 0), (\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'\mathbf{Y} = \begin{pmatrix} -.018 \\ 2.053 \\ -3.447 \end{pmatrix} \text{ and}$$

$$SSE = .202.$$

10 pts

b) **Is there definitive evidence** that a simpler model $y \approx \beta_0 + \beta_1 x \quad \forall x$ is inadequate here? **Explain.**

10 pts

c) Tomorrow a total of 3 new observations are to be drawn from this model at, respectively, $x = 1, 2,$ and 3 . Call these $y_1^*, y_2^*,$ and y_3^* . The quantity $(y_3^* - y_2^*) - (y_2^* - y_1^*) = y_3^* - 2y_2^* + y_1^*$ is an empirical measure of change in slope of mean y as a function of x at $x = 2$ based on these new observations. **Provide** 95% two-sided **prediction limits** for this quantity. (**Plug in completely**, but you need not do arithmetic.)

10 pts

d) **Find the value and degrees of freedom** for a t statistic for testing $H_0 : \mu_{y|x=1} = \mu_{y|x=5}$ (the hypothesis that the mean responses are the same for $x = 1$ and $x = 5$).

$T =$ _____

$df =$ _____

10 pts

e) **Write out** (plug in completely so that your implied answer is numerical, but you need not do the arithmetic) **a test statistic** that you could use to test the hypothesis that $H_0 : \mu_{y|x=1} = \mu_{y|x=5} = 1$. **Say** exactly what null distribution you would use.

10 pts

f) It is possible to compute $\mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'$ both for the full model specified at the beginning of this problem and for a model with \mathbf{X} matrix consisting of only the first two columns of the original one. The diagonal entries of these two matrices are in the table below.

i	1	2	3	4	5	6
x_i	0	1	2	3	4	5
diagonal entry of $\mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'$ for the original \mathbf{X}	.825	.298	.614	.272	.298	.693
diagonal entry of $\mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'$ for the reduced \mathbf{X}	.524	.295	.181	.181	.295	.524

Compare the two patterns above and **say** why (in the context provided at the beginning of this problem) they "make sense."

10 pts

2. Suppose that \mathbf{Y} is $MVN_n(\boldsymbol{\mu}, \sigma^2 \mathbf{I})$ and that \mathbf{A}, \mathbf{B} , and \mathbf{C} are symmetric $n \times n$ matrices with $\mathbf{AB} = \mathbf{0}$, $\mathbf{AC} = \mathbf{0}$, and $\mathbf{BC} = \mathbf{0}$. **Argue carefully** that the three random variables $\mathbf{Y}'\mathbf{A}\mathbf{Y}$, $\mathbf{Y}'\mathbf{B}\mathbf{Y}$, and $\mathbf{Y}'\mathbf{C}\mathbf{Y}$ are *jointly* independent.

10 pts

3. Suppose that y_{11} and y_{12} are independent $N(\mu_1, \eta)$ variables independent of y_{21} and y_{22} that are independent $N(\mu_2, 4\eta)$ variables. (The η and 4η are variances.) **What is the BLUE** of $\mu_1 - \mu_2$? **Explain** carefully.

5pts

4. a) For any non-zero $\mathbf{w} \in \mathfrak{R}^n$ the set of multiples of \mathbf{w} , namely $\{c\mathbf{w} \mid c \in \mathfrak{R}\}$, is a 1-dimensional subspace of \mathfrak{R}^n . We might call this subspace $C(\mathbf{w})$. Consider the operation of perpendicular projection onto $C(\mathbf{w})$, accomplished using the $n \times n$ projection matrix $\mathbf{P}_{\mathbf{w}}$. **Argue carefully** that for any $\mathbf{v} \in \mathfrak{R}^n$,

$$\mathbf{P}_{\mathbf{w}} \mathbf{v} = \left(\frac{\mathbf{v}'\mathbf{w}}{\mathbf{w}'\mathbf{w}} \right) \mathbf{w}$$

(Note that $\mathbf{P}_{\mathbf{w}} \mathbf{v} = c\mathbf{w}$ for some $c \in \mathfrak{R}$, and consider $c\mathbf{w}'\mathbf{w}$.)

5pts

b) In the regression context from lecture, let $\mathbf{X} = (\mathbf{1} \mid \mathbf{x}_1 \mid \mathbf{x}_2 \mid \cdots \mid \mathbf{x}_{r-1} \mid \mathbf{x}_r)$ and $\mathbf{X}_{r-1} = (\mathbf{1} \mid \mathbf{x}_1 \mid \mathbf{x}_2 \mid \cdots \mid \mathbf{x}_{r-1})$. Further, let

$$\mathbf{z}_r = \mathbf{x}_r - \mathbf{P}_{\mathbf{X}_{r-1}} \mathbf{x}_r = (\mathbf{I} - \mathbf{P}_{\mathbf{X}_{r-1}}) \mathbf{x}_r$$

Argue carefully that for any $\mathbf{v} \in C(\mathbf{X}_{r-1})$, $\mathbf{v} \perp \mathbf{z}_r$.

5pts

c) As a matter of fact, $\mathbf{P}_X - \mathbf{P}_{X_{r-1}} = \mathbf{P}_{z_r}$. **Argue carefully** here that $\mathbf{P}_X - \mathbf{P}_{X_{r-1}}$ is symmetric and idempotent, and that $(\mathbf{P}_X - \mathbf{P}_{X_{r-1}})\mathbf{v} = \mathbf{v}$ for any $\mathbf{v} \in C(\mathbf{z}_r)$.

5pts

d) Using the facts in a)-c) **argue carefully** that

$$\hat{\mathbf{Y}} = \mathbf{P}_{X_{r-1}} \mathbf{Y} + \left(\frac{\mathbf{e}'_{r-1} \mathbf{z}_r}{\mathbf{z}'_r \mathbf{z}_r} \right) \mathbf{z}_r$$

for $\mathbf{e}_{r-1} = (\mathbf{I} - \mathbf{P}_{X_{r-1}})\mathbf{Y}$. Then **say why** it is clear that the multiplier of \mathbf{z}_r here is b_r^{OLS} , the ordinary least squares estimate of the regression coefficient β_r in the full original regression. **What interpretation** does this development provide for b_r^{OLS} ?